## Stochastic Structural Dynamics

Lecture-21

## **Markov Vector Approach-1**

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#### Recall

# **Markov Property**

•Referes to a property displayed by conditional PDF-s of random processes.

A scalar random process X(t) is said to possess

Markov property if

$$\begin{aligned} &P\left[X\left(t_{n}\right) \leq x_{n} \mid X\left(t_{n-1}\right) \leq x_{n-1}, X\left(t_{n-2}\right) \leq x_{n-2}, \cdots, X\left(t_{1}\right) \leq x_{1}\right] \\ &= P\left[X\left(t_{n}\right) \leq x_{n} \mid X\left(t_{n-1}\right) \leq x_{n-1}\right] \end{aligned}$$

for any *n* and any choice of  $0 < t_1 < t_2 < \cdots < t_n$ .

# **Diffusion process:**

A Markov process with continuous state and continuous parameter

## Markov chain:

A markov process with discrete parameter

- •Mathematically, it is easier to deal with Markov chains.
- •In vibration problems, however, we are interested in diffusion processes.

# **Example**

Let  $\{\Theta_i\}_{i=1}^n$  be an iid sequence of random variables. Define

$$S_n = \sum_{i=1}^n \Theta_i$$

Define 
$$S_n = \sum_{i=1}^n \Theta_i$$
 
$$S_n = \Theta_n + \sum_{i=1}^{n-1} \Theta_i = \Theta_n + S_{n-1}$$
 
$$S_n \text{ depends only upon } S_{n-1} \text{ are } p(s_n \mid s_{n-1}) \text{ is not a function } 0$$

 $S_n$  depends only upon  $S_{n-1}$  and not upon  $\left\{S_i\right\}_{i=1}^{n-2}$   $p\left(s_n \mid s_{n-1}\right)$  is not a function of  $\left\{S_i = s_i\right\}_{i=1}^{n-2}$ 

$$p(s_n | s_{n-1})$$
 is not a function of  $\{S_i = s_i\}_{i=1}^{n-2}$ 

$$\Rightarrow S_n$$
 is Markov.  
Also,  $p_{S_n}(s \mid S_{n-1} = u) = p_{\Theta_n}(s - u)$ 

# Simple random walk, absorbing and reflecting boundaries

The model  $S_n = S_{n-1} + \Theta_n$ ;  $n = 1, 2, \dots$ 

represents the model for a simple random walk.

## Example 1

Consider an insurance company which starts business with an initial capital of  $X_0$  at time t = 0. Let

 $Y_1, Y_2, \dots, Y_n =$ premiums received

 $W_1, W_2, \dots, W_n = \text{claims paid}$ 

Capital at time n is given by

$$X_n = X_0 + (Y_1 - W_1) + (Y_2 - W_2) + \dots + (Y_n - W_n)$$

At any n, if  $X_n < 0$ , the company is ruined and it

can no longer perform.  $X_n = 0$  is called an absorbing barrier.

# Example 2 Gambler's ruin

Two gamblers A and B with capitals *a* and *b* play a game *n* number of times.

Game ends when

A wins all the capital b from B

or A loses all his capital a to B

 $X_n$  = capital gain of A at the end of  $n^{th}$  game.

X = -a and X = b are absorbing barriers.

This is an example of random walk with

two absorbing barriers.

# Example 3 Daily water level in a dam

b = capacity of dam

 $X_n$  = volume of water in the dam on the  $n^{\text{th}}$  day.

 $Y_n = \text{inflow on } n^{\text{th}} \text{ day.}$ 

 $X_n$  executes a random walk in the interval 0 to b

 $X_n = 0 \& X_n = b$  are the two reflecting barriers.

## **Types of questions**

# •Insurance company

What is the probability of ruin of the insurance company for a given  $X_0$ ?

# •Two gamblers

What is the probability of ruin of A and B?

# Dam filling

- •What is the long-term equillibrium probability of water level in the dam?
- •What is the PDF of empty periods and non-empty periods?

#### Example

Verify if an iid sequence of random variables form a Markov process.

• 
$$p(x_n | x_{n-1}, x_{n-2}, \dots, x_1) = p(x_n | x_{n-1}) = p(x_n)$$
 [Yes]

## Example

Let X(t) = A + Bt where A and B are iid random variables.

Is X(t) Markov? [No]

#### Example

Let A, B, C and D be random variables. Consider

Investigate the conditions under which X(t) is Markov.

#### Complete specification of a Markov process

$$\bullet P(x_1;t_1) = P[X(t_1) \le x_1]$$

$$\bullet P(x_1, x_2; t_1, t_2) = P[X(t_2) \le x_2 \mid X(t_1) \le x_1] P[X(t_1) \le x_1]$$

$$\bullet P(x_1, x_2, x_3; t_1, t_2, t_3)$$

$$= P\left[X\left(t_3\right) \le x_3 \mid X\left(t_2\right) \le x_2, X\left(t_1\right) \le x_1\right]$$

$$P\left[X\left(t_{2}\right) \leq x_{2} \mid X\left(t_{1}\right) \leq x_{1}\right] P\left[X\left(t_{1}\right) \leq x_{1}\right]$$

$$= P \left[ X \left( t_3 \right) \le x_3 \mid X \left( t_2 \right) \le x_2 \right] \smile$$

$$P[X(t_1) \le x_1 \mid X(t_1) \le x_1]P[X(t_1) \le x_1]$$

$$= \prod_{\nu=2}^{3} P \left[ X \left( t_{\nu} \right) \le x_{\nu} \mid X \left( t_{\nu-1} \right) \le x_{\nu-1} \right] P \left[ X \left( t_{1} \right) \le x_{1} \right]$$

$$\bullet P(x_1, x_2, \dots, x_n; t_1, t_2, \dots, t_n) =$$

$$\prod_{\nu=2}^{n} P\left[X\left(t_{\nu}\right) \leq x_{\nu} \mid X\left(t_{\nu-1}\right) \leq x_{\nu-1}\right] P\left[X\left(t_{1}\right) \leq x_{1}\right]$$

$$\bullet P \Big[ X \Big( t_{\nu} \Big) \le x_{\nu} \mid X \Big( t_{\nu-1} \Big) \le x_{\nu-1} \Big]$$

$$\&P[X(t_1) \le x_1] \forall n \& \{t_v\}_{v=1}^n$$

Complete specification
$$\bullet P \Big[ X \Big( t_{\nu} \Big) \le x_{\nu} \mid X \Big( t_{\nu-1} \Big) \le x_{\nu-1} \Big] \\
\bullet P \Big[ X \Big( t_{1} \Big) \le x_{1} \Big] \forall n \& \Big\{ t_{\nu} \Big\}_{\nu=1}^{n} \\
\bullet P \Big[ X \Big( t_{\nu} \Big) \le x_{\nu}, X \Big( t_{\nu-1} \Big) \le x_{\nu-1} \Big] \forall n \& \Big\{ t_{\nu} \Big\}_{\nu=1}^{n}$$

$$\bullet P(x_1, x_2, \dots, x_n; t_1, t_2, \dots, t_n)$$

$$= \prod_{\nu=2}^{n} P\left[X\left(t_{\nu}\right) \le x_{\nu} \mid X\left(t_{\nu-1}\right) \le x_{\nu-1}\right] P\left[X\left(t_{1}\right) \le x_{1}\right]$$

•Transistional PDF of 
$$X(t)$$
 [TPDF] 
$$P[X(t_{\nu}) \le x_{\nu} \mid X(t_{\nu-1}) \le x_{\nu-1}]$$
•TPDF represents the evolution mechanism.

- $\bullet P \Big[ X \Big( t_{n+1} \Big) \le x_{n+1} \mid X \Big( t_n \Big) \le x_n, X \Big( t_{n-1} \Big) \le x_{n-1} \Big] \\
  = P \Big[ X \Big( t_{n+1} \Big) \le x_{n+1} \mid X \Big( t_n \Big) \le x_n \Big].$

$$= P \left[ X \left( t_{n+1} \right) \le x_{n+1} \mid X \left( t_n \right) \le x_n \right].$$

$$\bullet P \Big[ X \Big( t_{n+1} \Big) \le x_{n+1} \mid X \Big( t_n \Big) \le x_n, X \Big( t_{n-1} \Big) \le x_{n-1} \Big] \\
= P \Big[ X \Big( t_{n+1} \Big) \le x_{n+1} \mid X \Big( t_n \Big) \le x_n \Big] \\
n+1 = \mathbf{Tomorrow}$$

n-1 = Yesterday

Tomorrow's happenings depend on the what is happening today and and not on what happened yesterday!

#### Recall: Model for dynamical systems using ODE - s

$$\frac{dy}{dt} = h [y(t), t]; y(0) = y_0$$
Let  $t_2 > t_1$ .
$$y(t_2) = q[t_2, y(t_1), t_1]$$

$$y(t_2) = q[t_2, y(t_1), t_1]$$

That is, solution at  $t_2$  depends upon  $y(t_1)$  and not on  $y(\tau)$   $(0 < \tau < t_1)$ .

Markov property can be viewed as the stochastic analog of the above property of ODE-s in deterministic systems

•Transistional prbability density function [tpdf]

$$p(x_{v};t_{v} | x_{v-1};t_{v-1}) = \frac{\partial}{\partial x_{v}} P[X(t_{v}) \leq x_{v} | X(t_{v-1}) = x_{v-1}]$$

$$p(x_{1},x_{2},\dots,x_{n};t_{1},t_{2},\dots,t_{n}) = p(x_{n};t_{n} | x_{n-1};t_{n-1},x_{n-2};t_{n-2},\dots,x_{1};t_{1})$$

$$p(x_{n-1};t_{n-1} | x_{n-2};t_{n-2},\dots,x_{1};t_{1})$$

$$\dots p(x_{2};t_{2} | x_{1};t_{1})p(x_{1};t_{1})$$

$$= p(x_{n};t_{n} | x_{n-1};t_{n-1})p(x_{n-1};t_{n-1} | x_{n-2};t_{n-2})$$

$$\dots p(x_{2};t_{2} | x_{1};t_{1})p(x_{1};t_{1})$$

$$= p(x_{1};t_{1})\prod_{v=2}^{n} p(x_{v};t_{v} | x_{v-1};t_{v-1})$$

#### Remark

Markov process is also completely specified in terms of

• 
$$p(x_{\nu}; t_{\nu}, x_{\nu-1}; t_{\nu-1}) \forall n \& \{t_{\nu}\}_{\nu=1}^{n}$$

$$\underbrace{p\left(x_{1}, x_{2}, \cdots, x_{n}; t_{1}, t_{2}, \cdots, t_{n}\right)}_{\text{Multi-dimensional jpdf}} = \underbrace{p(x_{1}; t_{1})}_{\text{Initial pdf}} \underbrace{\prod_{v=2}^{n} p\left(x_{v}; t_{v} \mid x_{v-1}; t_{v-1}\right)}_{\text{Product of transistional pdfs}}$$

# Chapman - Kolmogorov - Smoluchowski Equation

$$t = t_{1} \qquad t = \tau$$

$$p(x_{2}, t_{2}; x_{1}, t_{1}) = p(x_{2}, t_{2} | x_{1}, t_{1}) p(x_{1}, t_{1})$$

$$= \int p(x_{2}, t_{2}; x, \tau; x_{1}, t_{1}) dx$$

$$= \int p(x_{2}, t_{2} | x, \tau; x_{1}, t_{1}) p(x, \tau | x_{1}, t_{1}) p(x_{1}, t_{1}) dx$$

$$\Rightarrow$$

$$p(x_{2}, t_{2} | x_{1}, t_{1}) = \int p(x_{2}, t_{2} | x, \tau; x_{1}, t_{1}) p(x, \tau | x_{1}, t_{1}) dx$$

$$= \int p(x_{2}, t_{2} | x, \tau) p(x, \tau | x_{1}, t_{1}) dx$$

# Consistency condition for the process to be Markov

$$p(x_{2}, t_{2} | x_{1}, t_{1}) = \int p(x_{2}, t_{2} | x, \tau) p(x, \tau | x_{1}, t_{1}) dx$$
for all  $t_{1} < \tau < t_{2}$ 

## **Question**

How to utilize this result in characterizing response of randomly driven dynamical systems?

## **Independent increment processes**

Let X(t) be a random process with continuous state and continuous parameter (time t).

Let  $t_1 < t_2 < \cdots < t_n$  be *n* time instants.

Define 
$$X(t_1, t_2) = X(t_2) - X(t_1)$$

 $X(t_1,t_2)$  is called the increment of X(t) on  $[t_1,t_2)$ .

If, for all  $t_1 < t_2 < \cdots < t_n$ , the increments  $X(t_1, t_2), X(t_2, t_3)$ ,

 $\cdots$ ,  $X(t_{n-1},t_n)$  form a sequence of independent random variables, then the process X(t) is said to be a process with independent increments.

#### Remark

Let X(t)  $(t \ge 0)$  be a random process with independent increments.

Define 
$$Y(t) = X(t) - X(0)$$
 with  $(t \ge 0)$ .

Y(t) is a process with independent increments and also

has the property 
$$P[Y(0)=0]=1$$
//

Also,  $Y(t_{n-1}, t_n) = X(t_{n-1}, t_n)$  and has the properties of incremets of X(t).

Without loss of generality, therefore,

it may be taken that P[X(0)=0]=1.

Processes with independent increments possess

Markovian property

Define 
$$Y(t) = X(t)$$
 if  $P[X(0) = 0] = 1$   
=  $X(t) - X(0)$  if  $P[X(0) = 0] \neq 1$ 

$$\Rightarrow Y(t_0) = 0$$

Consider 
$$t_n > t_{n-1} > t_{n-2} > \dots > t_1 > t_0 = 0$$

and the associated random variables  $Y(t_n), Y(t_{n-1}), \dots, Y(t_1)$ 

$$Y(t_n) = \sum_{j=1}^{n} \left[ Y(t_j) - Y(t_{j-1}) \right]$$

$$= Y(t_1) - Y(t_0) + Y(t_2) - Y(t_1) \cdots + Y(t_n) - Y(t_{n-1}) = Y(t_n)$$

Processes with independent increments possess

Denote 
$$Z_j = Y(t_j) - Y(t_{j-1})$$

 $\Rightarrow Z_j (j=1,2,\dots,n)$  are sequence of independent random variables.

$$Y(t_n) = \sum_{j=1}^n Z_j = Z_n + Y(t_{n-1})$$

$$\Rightarrow Y(t) \text{ is Markov.}$$

## Generalization: Markov property for vector random processes

Let X(t) be a vector random process with continuous state and continuous parameter (time t).

Let  $t_1 < t_2 < \dots < t_n$  be the *n* time instants.

This defines *n* vector random variables

$$X(t_1), X(t_2), \dots, X(t_n).$$

The vector process X(t) is said to possess Markov property if the  $n^{th}$  order conditional joint PDF

$$P\left[X\left(t_{n}\right) \leq x_{n} \mid X\left(t_{n-1}\right) \leq x_{n-1}, X\left(t_{n-2}\right) \leq x_{n-2}, \dots, X\left(t_{1}\right) \leq x_{1}\right]$$

$$= P\left[X\left(t_{n}\right) \leq x_{n} \mid X\left(t_{n-1}\right) \leq x_{n-1}\right]$$

for any *n* and any choice of  $t_1 < t_2 < \cdots < t_n$ .

The *m*-dimensional vector random process

$$\boldsymbol{X}(t) = \begin{bmatrix} X_1(t) & X_2(t) & \cdots & X_m(t) \end{bmatrix}^t$$

is said to be Markov if

$$P\left[\bigcap_{i=1}^{m} \left\{ X_{j}\left(t_{n}\right) \leq X_{j} \right\} \mid X\left(t_{n-1}\right) = y_{n-1}, X\left(t_{n-2}\right) = y_{n-2}, \cdots, X\left(t_{1}\right) = y_{1}\right]$$

$$= P\left[\bigcap_{i=1}^{m} \left\{X_{j}\left(t_{n}\right) \leq x_{j}\right\} \mid \boldsymbol{X}\left(t_{n-1}\right) = y_{n-1}\right] \forall t_{n} > t_{n-1} > \dots > t_{1}$$

$$\boldsymbol{TPDF} = P\left[\bigcap_{i=1}^{m} \left\{X_{j}\left(t_{n}\right) \leq x_{j}\right\} \mid \boldsymbol{X}\left(t_{n-1}\right) = y_{n-1}\right]$$

$$TPDF = P \left[ \bigcap_{i=1}^{m} \left\{ X_{j} \left( t_{n} \right) \leq X_{j} \right\} \mid X \left( t_{n-1} \right) = y_{n-1} \right]$$

$$\widehat{p(\mathbf{x};t \mid \mathbf{x}_{0}, t_{0})} = \frac{\partial}{\partial x_{1} \partial x_{2} \cdots \partial x_{m}} P \left[ \bigcap_{i=1}^{m} \left\{ X_{j}(t_{n}) \leq x_{j} \right\} \mid \mathbf{X}(t_{n-1}) = y_{n-1} \right]$$

#### **Remarks**

- •A component of a vector random process (which itself is a scalar random process) need not have Markov property
- •Different components of a Markov vector random process could be differentiable (in the mean square sense) to different levels.
- •If a vector random process has independent random increments (vectorially), the components of the random process need not have independent increments.

Consistency condition for a vector Markov process (CKS equation)

$$p(\tilde{x}_{3};t_{3} | \tilde{x}_{1};t_{1}) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \cdots \int_{-\infty}^{\infty} p(\tilde{x}_{3};t_{3} | \tilde{y};t_{2}) p(\tilde{y};t_{2} | \tilde{x}_{1};t_{1}) dy_{1} dy_{2} \cdots dy_{n}$$

## Generalization: higher order Markov property

Let X(t) be a scalar random process with continuous state and continuous parameter (time t).

Let  $t_1 < t_2 < \dots < t_n$  be *n* time instants.

This defines n random variables

$$X(t_1), X(t_2), \dots, X(t_n).$$

X(t) is said to possess  $2^{nd}$  order Markov property if

$$P\left[X\left(t_{n}\right) \leq \mathbf{X}_{n} \mid X\left(t_{n-1}\right) \leq \mathbf{X}_{n-1}, X\left(t_{n-2}\right) \leq \mathbf{X}_{n-2}, \cdots, X\left(t_{1}\right) \leq \mathbf{X}_{1}\right]$$

$$= P \left[ X(t_n) \le t_n \mid X(t_{n-1}) \le t_{n-1}, X(t_{n-2}) \le t_{n-2} \right]$$

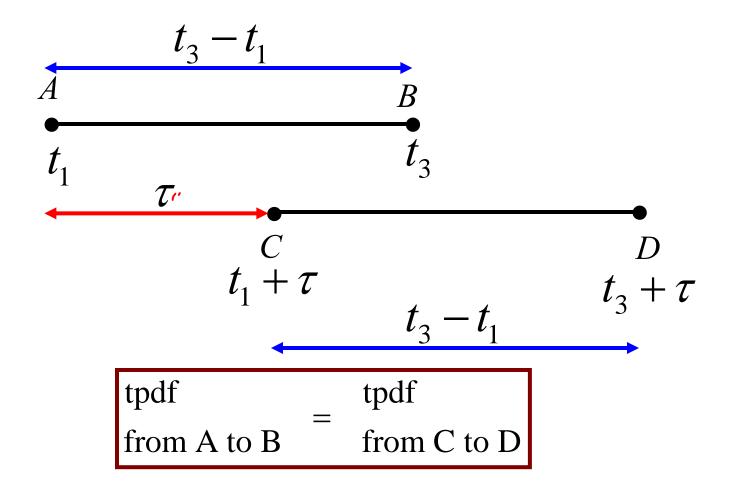
for any *n* and any choice of  $t_1 < t_2 < \cdots < t_n$ .

The idea can be generalized to 3<sup>rd</sup> and higher orders.

## Markov process with stationary increments

X(t) is said to be a Markov process with stationary increments if

$$p(x_3;t_3+\tau \mid x_1;t_1+\tau) = p(x_3;t_3 \mid x_1;t_1) \forall t_3 > t_1 \& \tau$$



#### Remarks

- $\Rightarrow$  tpdf can be written as  $p(x_3; \tau \mid x_1)$ ;

 $\tau = \text{transistion time} / 1$ 

- $\bullet \lim_{\tau \to \infty} p(x_3; \tau \mid x_1) \to p(x_3; \tau)$
- •Stationary Markov random process is completely specified in terms of the tpdf.
- $\bullet X(t)$  is stationary  $\Rightarrow X(t)$  has stationary increments.
- •X(t) has stationary in ements need not mean that X(t) is stationary.

#### **Kinetic equation**

Let X(t) be a scalar random process.

Let  $p_X(x;t)$  be the first order pdf of X(t).

Can we derive a differential equation satisfied by  $p_X(x;t)$ ?

Here x and t would be the independent variables and  $p_X(x;t)$ 

would be the dependent variable. Thus the differential equation

that we are looking for would be a partial differential equation.

Furthermore, can we also derive the PDE governing

the *n*-dimensional joint pdf of X(t)?

Notation:  $p_{X}(x;t) = p(x;t)$ 

Reference: T T Soong, 1973, Random differential equations in science and engineering, Academic Press, NY.

Consider the random variables  $X(t) \& X(t + \Delta t)$ 

$$p(x,t+\Delta t) = \int_{-\infty}^{\infty} p(x,x';t,t+\Delta t) dx'$$

$$= \int_{-\infty}^{\infty} p(x; t + \Delta t \mid x'; t) \underline{p(x'; t)} dx' \cdots (1)$$

Define 
$$\Delta X(t) = X(t + \Delta t) - X(t)$$
.

$$\langle \exp[iu\Delta X(t)|X(t)=x']\rangle$$
 = Conditional characteristic

function of  $\Delta X(t)$  given X(t) = x'.

Denote 
$$\Phi(u, t + \Delta t \mid \underline{X(t)} = x') = \langle \exp[iu\Delta X(t) \mid X(t) = x'] \rangle$$
.

$$\Phi(u, t + \Delta t \mid X(t) = x') = \int_{-\infty}^{\infty} \exp(iu\Delta x) \underline{p(x; t + \Delta t \mid X(t) = x')} dx \cdots (2)$$
with  $\Delta X(t) = X(t + \Delta t) - X(t)$  such that  $\Delta x = x - x' \cdots (3)$ .

$$\Phi\left(u,t+\Delta t\mid X(t)=x'\right) = \int_{-\infty}^{\infty} \exp\left(iu\Delta x\right) p\left(x;t+\Delta t\mid X(t)=x'\right) dx$$

$$\Rightarrow p\left(x;t+\Delta t\mid X(t)=x'\right)$$

$$= \frac{1}{2\pi} \int_{-\infty}^{\infty} \exp\left(-iu\Delta x\right) \Phi\left(u,t+\Delta t\mid X(t)=x'\right) du \cdots (4)$$

$$= \frac{1}{2\pi} \int_{-\infty}^{\infty} \exp\left(-iu\Delta x\right) \Phi\left(u,t+\Delta t\mid X(t)=x'\right) du \cdots (4)$$

$$\Rightarrow \Phi\left(u,t+\Delta t\mid X(t)=x'\right) = \Phi\left(0+u,t+\Delta t\mid X(t)=x'\right)$$

$$= \sum_{n=0}^{\infty} \frac{u^n}{n!} \frac{d^n}{du^n} \Phi\left(u,t+\Delta t\mid X(t)=x'\right)\Big|_{u=0}$$

$$= \sum_{n=0}^{\infty} \frac{(iu)^n}{n!} \left\langle \left[\Delta X^n(t)\mid X(t)=x'\right] \right\rangle$$

## Recall: Characteristic function of a random variable

Definition: 
$$\varphi_X(\omega) = \langle \exp(i\omega X) \rangle = \int_{-\infty}^{\infty} \exp(i\omega x) p_X(x) dx$$
  
Here  $\omega$  is real valued. Thus,  $\varphi_X(\omega)$  is the Fourier transform of the pdf.

It can be shown that 
$$\frac{1}{i^n} \frac{d^n \varphi_X}{d\omega^n} \bigg|_{\omega=0} = \langle X^n \rangle$$
.

By using inverse Fourier transform, it follows that

$$p_X(x) = \int_{-\infty}^{\infty} \varphi_X(\omega) \exp(-i\omega x) dx.$$

#### Characteristic function of a random process

$$\alpha_{j_1,j_2,\cdots,j_n}\left(t_1,t_2,\cdots,t_n\right) = \left\langle X^{j_1}\left(t_1\right)X^{j_2}\left(t_2\right)\cdots X^{j_n}\left(t_n\right)\right\rangle$$

$$\frac{\partial^m}{\partial u_1^{j_1} \partial u_2^{j_2} \cdots \partial u_n^{j_n}} \phi_X \left( u_1; t_1, u_2; t_2, \cdots, u_n; t_n \right) \Big|_{\underbrace{\left\{ u_j \right\}_{j=1}^n = 0}} = i^m \alpha_{j_1, j_2, \cdots, j_n} \left( t_1, t_2, \cdots, t_n \right)$$

such that  $m = j_1 + j_2 + \cdots + j_n$ .

$$\Rightarrow p(x;t+\Delta t \mid X(t)=x')=$$

$$\Rightarrow p(x;t+\Delta t \mid X(t) = x') = \frac{1}{2\pi} \int_{-\infty}^{\infty} \exp(-iu\Delta x) \left\{ \sum_{n=0}^{\infty} \frac{(iu)^n}{n!} \left\langle \left[ \Delta X^n(t) \mid X(t) = x' \right] \right\rangle \right\} du$$

$$p(x;t+\Delta t \mid X(t)=x') = \sum_{n=0}^{\infty} \frac{i^n}{2\pi} \frac{1}{n!} a_n(x',t) \int_{-\infty}^{\infty} u^n \exp(-iu\Delta x) du$$

with

$$a_{n}(x',t) = \left\langle \left[ \Delta X^{n}(t) \mid \underline{X}(t) = \underline{x'} \right] \right\rangle$$

$$= \left\langle \left[ \left\{ X(t + \Delta t) - x(t) \right\}^{n} \mid X(t) = \underline{x'} \right] \right\rangle \cdots (5)$$

#### = Incremental moments

$$p(x;t+\Delta t \mid X(t)=x') = \sum_{n=0}^{\infty} \frac{a_n(x',t)}{n!} \frac{1}{2\pi} \int_{-\infty}^{\infty} (iu)^n \exp(-iu\Delta x) du$$

$$p(x;t+\Delta t \mid X(t)=x') = \sum_{n=0}^{\infty} \frac{a_n(x',t)}{n!} \frac{1}{2\pi} \int_{-\infty}^{\infty} (iu)^n \exp(-iu\Delta x) du \cdots (6)$$

#### Recall

$$\int_{-\infty}^{\infty} \delta(x) \exp(iux) dx = 1$$

$$\Rightarrow \delta(x) = \frac{1}{2\pi} \int_{-\infty}^{\infty} \exp(-iux) du$$

$$\Rightarrow \frac{d^{n}}{dx^{n}} \delta(x) = \frac{1}{2\pi} \int_{-\infty}^{\infty} (-1)^{n} (iu)^{n} \exp(-iux) du$$

$$\Rightarrow \delta(x) = \frac{1}{2\pi} \int_{-\infty}^{\infty} \exp(-iux) du$$

$$\Rightarrow \frac{d^n}{dx^n} \delta(x) = \frac{1}{2\pi} \int_{-\infty}^{\infty} (-1)^n (iu)^n \exp(-iux) du / D$$

$$p(x; t + \Delta t \mid X(t) = x') = \sum_{n=0}^{\infty} \frac{(-1)^n}{n!} a_n(x', t) \frac{d^n}{dx^n} \delta(\Delta x) \cdots (7)$$

#### Substitute equation 7 in 1

$$p(x;t+\Delta t \mid X(t) = x') = \sum_{n=0}^{\infty} \frac{(-1)^n}{n!} a_n(x',t) \frac{d^n}{dx^n} \delta(\Delta x) \cdots (7)$$

$$p(x,t+\Delta t) = \int_{-\infty}^{\infty} p(x;t+\Delta t \mid x';t) p(x';t) dx' \cdots (1)$$

$$p(x,t+\Delta t) = \int_{-\infty}^{\infty} p(x;t+\Delta t \mid x';t) p(x';t) dx' \cdots (1)$$

$$\Rightarrow p(x,t+\Delta t) = \int_{-\infty}^{\infty} \sum_{n=0}^{\infty} \frac{(-1)^n}{n!} a_n(x',t) \frac{d^n}{dx^n} \delta(\Delta x) p(x';t) dx'$$

$$= \sum_{n=1}^{\infty} \frac{(-1)^n}{n!} \frac{d^n}{dx^n} \int_{-\infty}^{\infty} \delta(\Delta x) a_n(x',t) p(x';t) dx'$$

$$\operatorname{Recall} \frac{d^n}{dx^n} \int_{-\infty}^{\infty} \delta(x-a) f(x) dx = \frac{d^n f(a)}{dx^n}$$

Recall 
$$\frac{d^n}{dx^n} \int_{-\infty}^{\infty} \delta(x-a) f(x) dx = \frac{d^n f(a)}{dx^n}$$

$$p(x,t+\Delta t) = \sum_{n=0}^{\infty} \frac{\left(-1\right)^n}{n!} \frac{d^n}{dx^n} \left[ a_n(x,t) p(x;t) \right]$$

$$p(x,t+\Delta t) = \sum_{n=0}^{\infty} \frac{(-1)^n}{n!} \frac{d^n}{dx^n} \left[ a_n(x,t) p(x;t) \right]$$

$$= a_0(x,t) p(x;t) + \sum_{n=1}^{\infty} \frac{(-1)^n}{n!} \frac{d^n}{dx^n} \left[ a_n(x,t) p(x;t) \right] \dots (8)$$

$$a_0(x,t) = \left\langle \Delta X^0(t) | X(t) = x' \right\rangle = 1$$

$$\Rightarrow p(x,t+\Delta t) - p(x;t) = \sum_{n=1}^{\infty} \frac{(-1)^n}{n!} \frac{d^n}{dx^n} \left[ a_n(x,t) p(x;t) \right]$$

$$\frac{\partial p}{\partial t} = \lim_{\Delta t \to 0} \frac{1}{\Delta t} \sum_{n=1}^{\infty} \frac{(-1)^n}{n!} \frac{d^n}{dx^n} \left[ a_n(x,t) p(x;t) \right]$$

$$= \sum_{n=1}^{\infty} \frac{(-1)^n}{n!} \lim_{\Delta t \to 0} \frac{1}{\Delta t} \frac{d^n}{dx^n} \left[ a_n(x,t) p(x;t) \right]$$

$$= \sum_{n=1}^{\infty} \frac{(-1)^n}{n!} \frac{d^n}{dx^n} \left[ \alpha_n(x,t) p(x;t) \right]$$

$$\frac{\partial p}{\partial t} = \sum_{n=1}^{\infty} \frac{\left(-1\right)^n}{n!} \frac{\partial^n}{\partial x^n} \left[\alpha_n(x,t) p(x;t)\right]$$

$$\alpha_{n}(x,t) = \lim_{\Delta t \to 0} \frac{1}{\Delta t} \left\langle \left[ X(t + \Delta t) - X(t) \right]^{n} \mid X(t) = x \right\rangle; n = 1, 2, \dots$$

= Derivative moments

$$\lambda(x,t) = -\sum_{n=1}^{\infty} \frac{\left(-1\right)^n}{n!} \frac{\partial^{n-1}}{\partial x^{n-1}} \left[\alpha_n(x,t) p(x;t)\right]$$

$$\frac{\partial p}{\partial t} + \frac{\partial \lambda}{\partial x} = 0$$

Equation of conservation of probability

Similar to equation of continuity in fluid mechanics.

- Diffusion equation
- • $\lambda(x,t)$ =amount of probability crossing x in unit time

#### **Remarks**

•The equation  $\frac{\partial p}{\partial t} + \frac{\partial \lambda}{\partial t} = 0$  has infinite order with

respect to the spatial coordinate.

Therefore, its application is severely limited.

#### •Theorem

R F Pawula, 1967, IEEE Trans.

Information theory, IT-13, 33-41

If the derivative moment  $\alpha_n(x,t)$  exists for all n and is zero for some even n, then  $\alpha_n(x,t) = 0 \forall n \geq 3$ .

# **Implication:**

There exist two cases of equation

$$\frac{\partial p}{\partial t} + \frac{\partial \lambda}{\partial t} = 0.$$
 The one in which

the order of the equaiton in x is infinite and the other in which the order is 2 or less. We would be interested in applying the kinetic equation for the case in which  $\alpha_n(x,t) = 0 \forall n \geq 3$ .

#### Proof of the statement:

If the derivative moment  $\alpha_n(x,t)$  exists for all n and is zero for some even n, then  $\alpha_n(x,t) = 0 \forall n \geq 3$ .

Let  $n \ge 3$  and let n be odd.

$$\alpha_{n}(x,t) = \lim_{\Delta t \to 0} \frac{1}{\Delta t} \left\langle \left[ X(t + \Delta t) - X(t) \right]^{n} | X(t) = x \right\rangle$$

$$= \lim_{\Delta t \to 0} \frac{1}{\Delta t} \left\langle \left[ X(t + \Delta t) - X(t) \right]^{\frac{n-1}{2}} \left[ X(t + \Delta t) - X(t) \right]^{\frac{n+1}{2}} | X(t) = x \right\rangle$$

Apply the Schwarz inequality

$$\alpha_{n}^{2}(x,t) \leq \lim_{\Delta t \to 0} \frac{1}{\Delta t^{2}} \left\langle \left[ X(t + \Delta t) - X(t) \right]^{n-1} | X(t) = x \right\rangle$$

$$\left\langle \left[ X(t + \Delta t) - X(t) \right]^{n+1} | X(t) = x \right\rangle.$$

$$\Rightarrow \alpha_{n}^{2}(x,t) \leq \alpha_{n-1}(x,t) \alpha_{n+1}(x,t) \forall n \geq 3 \& n \text{ odd} \cdots (1)$$

$$\left\langle \left[ X(t+\Delta t) - X(t) \right]^{n+1} | X(t) = x \right\rangle$$

$$\Rightarrow \alpha_n^2(x,t) \le \alpha_{n-1}(x,t)\alpha_{n+1}(x,t) \forall n \ge 3 \& n \text{ odd} \cdots (1)$$

Similarly it can be shown that

$$\alpha_n^2(x,t) \le \alpha_{n-2}(x,t)\alpha_{n+2}(x,t) \forall n \ge 4 \& n \text{ even} \cdots (2)$$

Let 
$$\alpha_r(x,t) = 0$$
 where  $r$  is an even integer.  
Let  $n = r - 1, r + 1$ , and using  $\alpha_n^2(x,t) \le \alpha_{n-1}(x,t) \alpha_{n+1}(x,t) \forall n \ge 3 \& n$  odd

$$\alpha_{r-1}^{2}(x,t) \le \alpha_{r-2}(x,t)\alpha_{r}(x,t) \forall r \ge 4$$

$$\alpha_{r-1}^{2}(x,t) \le \alpha_{r-2}(x,t)\alpha_{r}(x,t) \forall r \ge 4$$

$$\alpha_{r+1}^{2}(x,t) \le \alpha_{r}(x,t)\alpha_{r+2}(x,t) \forall r \ge 2$$

Let n = r - 2, r + 2, and using  $\alpha_n^2(x,t) \le \alpha_{n-2}(x,t)\alpha_{n+2}(x,t) \forall n \ge 4 \& n \text{ even}$ we get  $\alpha_{r-2}^2(x,t) \le \alpha_{r-4}(x,t)\alpha_r(x,t) \forall r \ge 6$   $\alpha_{r+2}^2(x,t) \le \alpha_r(x,t)\alpha_{r+4}(x,t) \forall r \ge 2$ 

$$\alpha_{r-1}^{2} \leq \alpha_{r-2}\alpha_{r} \forall r \geq 4$$

$$\alpha_{r+1}^{2} \leq \alpha_{r}\alpha_{r+2} \forall r \geq 2$$

$$\alpha_{r-2}^{2} \leq \alpha_{r-4}\alpha_{r} \forall r \geq 6$$

$$\alpha_{r+2}^{2} \leq \alpha_{r}\alpha_{r+4} \forall r \geq 2$$
Illustration:

Let  $\alpha_3 = 0$  and  $\alpha_n$  exist for all n.

$$\alpha_{r+1}^{2} \le \alpha_{r} \alpha_{r+2} \forall r \ge 2 \Rightarrow \alpha_{4}^{2} = 0$$

$$\alpha_{r+2}^{2} \le \alpha_{r} \alpha_{r+4} \forall r \ge 2 \Rightarrow \alpha_{5}^{2} = 0$$

$$\Rightarrow \alpha_{r} = 0 \forall r > 3$$

$$\alpha_{r+2}^2 \le \alpha_r \alpha_{r+4} \forall r \ge 2 \Longrightarrow \alpha_5^2 = 0$$

$$\Rightarrow \alpha_r = 0 \forall r > 3$$

$$\alpha_{r-1}^{2} \le \alpha_{r-2}\alpha_{r} \forall r \ge 4$$

$$\alpha_{r+1}^{2} \le \alpha_{r}\alpha_{r+2} \forall r \ge 2$$

$$\alpha_{r-2}^{2} \le \alpha_{r-4}\alpha_{r} \forall r \ge 6$$

$$\alpha_{r+2}^{2} \le \alpha_{r}\alpha_{r+4} \forall r \ge 2$$

Let  $\alpha_r = 0$  and  $\alpha_n$  exist for all n.

$$\alpha_{r+2}^{2} \le \alpha_{r}\alpha_{r+4} \forall r \ge 2 \& \alpha_{r+1}^{2} \le \alpha_{r}\alpha_{r+2} \forall r \ge 2 \Rightarrow \alpha_{n} = 0 \forall n > r$$

$$\alpha_{r-2}^{2} \le \alpha_{r-4}\alpha_{r} \forall r \ge 6 \& \alpha_{r-1}^{2} \le \alpha_{r-2}\alpha_{r} \forall r \ge 4 \Rightarrow \alpha_{n} = 0 \forall n < r \& n \ge 3$$
OFD

#### Remarks

•For the case  $\alpha_n(x,t) = 0 \forall n \ge 3$ , the kinetic equation takes the form

$$\frac{\partial p}{\partial t} = -\frac{\partial}{\partial x} \left[ \alpha_1(x,t) p(x;t) \right] + \frac{1}{2} \frac{\partial^2}{\partial x^2} \left[ \alpha_2(x,t) p(x;t) \right]$$

This equation is known as the Fokker-Planck equation. Here

$$\lambda(x,t) = \alpha_1(x,t) p(x;t) - \frac{1}{2} \frac{\partial}{\partial x} \left[ \alpha_2(x,t) p(x;t) \right].$$

- •Initial condition:  $p(x;t=0) = p_0(x)$
- •Boundary conditions: different possibilities exist.

#### Boundary conditions

If X(t) takes values from  $-\infty$  to  $\infty$ , then the BCS are specified at  $\pm \infty$ . These boundaries are inaccessible.

$$\frac{\partial p}{\partial t} + \frac{\partial \lambda}{\partial t} = 0$$

$$\Rightarrow \int_{-\infty}^{\infty} \frac{\partial p}{\partial t} dx + \int_{-\infty}^{\infty} \frac{\partial \lambda}{\partial x} dx = 0$$

$$\Rightarrow \frac{\partial}{\partial t} \int_{-\infty}^{\infty} p(x;t) dx + \int_{-\infty}^{\infty} \frac{\partial \lambda}{\partial x} dx = 0$$

$$\int_{-\infty}^{\infty} p(x;t) dx = 1 \Rightarrow \int_{-\infty}^{\infty} \frac{\partial \lambda}{\partial x} dx = 0$$

$$\lambda(-\infty;t) = \lambda(\infty;t)$$

$$\lambda\left(-\infty;t\right) = \lambda\left(\infty;t\right)$$

A stronger condition:  $\lambda(-\infty;t) = \lambda(\infty;t) = 0$ 

Still stronger condition:  $p(\pm \infty; t) = 0$ 

If X(t) can take values only in the bounded region  $x_1 \le X(t) \le x_2$ , then the Fokker Planck equation is valid in this region with

$$\frac{\partial p}{\partial t} + \frac{\partial \lambda}{\partial t} = 0 \Rightarrow \frac{\partial}{\partial t} \int_{x_1}^{x_2} p(x;t) dx + \int_{x_1}^{x_2} \frac{\partial \lambda}{\partial x} dx = 0$$
$$\Rightarrow \lambda(x_1;t) = \lambda(x_2;t) / /$$

$$\Rightarrow \lambda(x_1;t) = \lambda(x_2;t)//$$

A stronger condition:  $\lambda(x_1;t) = \lambda(x_2;t) = 0$ 

Still stronger condition:  $p(x_1;t) = p(x_2;t) = 0$ 

$$\lambda(x_1;t) = \lambda(x_2;t) = 0$$
: Reflecting boundaries  $p(x_1;t) = p(x_2;t) = 0$ : Absorbing boundaries

$$p(x_1;t) = p(x_2;t) = 0$$
: Absorbing boundaries

### Remarks (continued)

•When  $\alpha_n(x,t)$  are independent of time, stationary solutions might exist such that

$$\frac{\partial p}{\partial t} = 0 \text{ and one gets the simplified equation}$$

$$-\frac{\partial}{\partial x} \left[ \alpha_1(x) p(x) \right] + \frac{1}{2} \frac{\partial^2}{\partial x^2} \left[ \alpha_2(x) p(x) \right] = 0$$
  
leading to the solution  $\lambda(x) = 0$ 

## Generalization of the kinetic equaiton for the case of a vector random process

Let  $X(t) = \{X_1(t), X_2(t), \dots, X_m(t)\}^t$  be a *m*-dimensional vector random process.

Consider the vector random variables  $X(t) \& X(t + \Delta t)$ 

$$p(\tilde{x}, t + \Delta t) = \int_{-\infty}^{\infty} p(\tilde{x}; t + \Delta t \mid \tilde{x}'; t) p(\tilde{x}'; t) d\tilde{x}'$$

$$p(\tilde{x},t+\Delta t) = \int_{-\infty}^{\infty} p(\tilde{x};t+\Delta t \mid \tilde{x}';t) p(\tilde{x}';t) d\tilde{x}'$$

$$\frac{\partial p(\tilde{x};t)}{\partial t} = \sum_{n_1,n_2,\dots,n_m=1}^{\infty} \left[ \prod_{j=1}^{m} \frac{(-1)^{n_j}}{(n_j)!} \frac{\partial}{\partial x_j^{n_j}} \right] \left[ \alpha_{n_1,n_2,\dots,n_m}(\tilde{x};t) p(\tilde{x};t) \right]$$

$$\alpha_{n_{1},n_{2},\cdots,n_{m}}\left(\tilde{x};t\right) = \lim_{\Delta t \to 0} \left\langle \prod_{j=1}^{m} \left[ X_{j}\left(t + \Delta t\right) - X_{j}\left(t\right) \right]^{n_{j}} |X\left(t\right) = \tilde{x} \right\rangle$$

## Next Lecture

- How to derive the Fokker Plank equation for response of dynamical systems driven by random excitations.
- How to solve them?