



INDIAN INSTITUTE OF SCIENCE

STOCHASTIC HYDROLOGY

Lecture -34

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Summary of the previous lecture

- Multivariate stochastic models
 - Cross correlation

$$r_{j,h}(k) = \frac{\sum_{i=1}^n (x_{j,i} - \bar{x}_j)(x_{h,i+k} - \bar{x}_h)}{(n-k)s_j s_h}$$

- Two site Markov model

$$x_{h,t} = \bar{x}_h + r_{j,h}(0) \frac{s_h}{s_j} (x_{j,t} - \bar{x}_j) + u_t s_h \sqrt{1 - r_{j,h}^2(0)}$$

Multivariate Stochastic models

Multisite Markov model:

- Multisite generation requires simultaneous generation of data at several sites while preserving the correlation between the data at various sites.
- Consider $x_{j,t}$

$$x_{j,t} = \frac{\left(x_{j,t}^d - \bar{x}_j \right)}{s_j}$$

Multivariate Stochastic models

$$X_{t+1} = \mu_x + \rho_1 (X_t - \mu_x) + u_{t+1} \sigma_x \sqrt{1 - \rho_1^2}$$

- The first order Markov model for site h is

$$x_{h,t+1} = \rho_h (1) x_{h,t} + \varepsilon_{h,t+1} \sqrt{1 - \rho_h^2 (1)}$$

$\mu = 0$ and $\sigma = 1$ because it is standardized data

- The first order Markov model for site j is

$$x_{j,t+1} = \rho_j (1) x_{j,t} + \varepsilon_{j,t+1} \sqrt{1 - \rho_j^2 (1)}$$

- The equations are written in matrix form



Multivariate Stochastic models

$$X_{t+1} = EX_t + G\varepsilon$$

where

X_t is a $p \times 1$ vector of standardized values of the variable generated at time t ,

E is a $p \times p$ diagonal matrix whose j^{th} diagonal element is $\rho_j(1)$,

G is a $p \times p$ diagonal matrix whose j^{th} diagonal element is $\sqrt{1 - \rho_j^2(1)}$

ε is a $p \times 1$ vector of random variates

Multivariate Stochastic models

- \mathcal{E} is defined to preserve the first order serial correlation (auto correlation) of the x_j 's and the lag zero cross-correlation between x_j and x_h .

- \mathcal{E} is made of elements that are $\varepsilon_{j,t+1}$; each $\varepsilon_{j,t+1}$ is independent of $x_{j,t}$; ε_j is $N(0, 1)$

- The cross correlation between ε_j and ε_h is $\rho_{j,h}^*(0)$,

$$\rho_{j,h}^*(0) = \frac{\{1 - \rho_j(1)\rho_h(1)\} \rho_{j,h}(0)}{\sqrt{\{1 - \rho_j^2(1)\}\{1 - \rho_h^2(1)\}}}$$

- $\rho_{j,h}^*(0)$ reproduces the desired $\rho_{j,h}(0)$, which is the lag zero cross correlation between x_j and x_h .

Multivariate Stochastic models

$$\boldsymbol{\varepsilon} = AD_{\lambda}^{1/2}e$$

where

$D_{\lambda}^{1/2}$ is a $p \times p$ diagonal matrix whose j^{th} diagonal element is the square root of the j^{th} largest eigenvalue of the $p \times p$ correlation matrix whose elements are $\rho_{j,h}^*(0)$

A is a $p \times p$ matrix consisting of eigenvectors of correlation matrix,

e is $p \times 1$ vector of independent observations from $N(0,1)$

Multivariate Stochastic models

- Matalas (1967) has given a multisite normal generation model that preserves the mean, variance, lag one serial correlation, lag one cross-correlation and lag zero cross-correlation.

$$X_{t+1} = AX_t + B\mathcal{E}_{t+1}$$

where

X_t and X_{t+1} are $p \times 1$ vectors representing standardized data corresponding to p sites at time steps t and $t+1$ resp.

Assumption is that the model is multivariate normal.

Ref.: Matalas, N.C. (1967) Mathematical assessment of synthetic hydrology, Water Resources Research 3(4):937-945

Multivariate Stochastic models

ε_{t+1} is $N(0,1)$; $p \times 1$ vector with ε_{t+1} independent of X_t .

A and B are coefficient matrices of size $p \times p$. B is assumed to be lower triangular matrix

$$X_{t+1} = \begin{bmatrix} x(1,t+1) \\ x(2,t+1) \\ \cdot \\ \cdot \\ x(i,t+1) \\ \cdot \\ \cdot \\ x(p,t+1) \end{bmatrix} \quad X_t = \begin{bmatrix} x(1,t) \\ x(2,t) \\ \cdot \\ \cdot \\ x(i,t) \\ \cdot \\ \cdot \\ x(p,t) \end{bmatrix} \quad \varepsilon_{t+1} = \begin{bmatrix} \varepsilon(1,t+1) \\ \varepsilon(2,t+1) \\ \cdot \\ \cdot \\ \varepsilon(i,t+1) \\ \cdot \\ \cdot \\ \varepsilon(p,t+1) \end{bmatrix}$$

Multivariate Stochastic models

- The scalar form is

$$x_{i,t+1} = \sum_{j=1}^p a_{i,j} x(j,t) + \sum_{j=1}^i b_{i,j} \varepsilon(i,t+1)$$

where

$a_{i,j}$ and $b_{i,j}$ denote the (i, j) th elements of the matrices A and B.

Multivariate Stochastic models

Coefficient matrices A and B:

- The expectation of $X_t X_t'$ is denoted by M_0

$$M_0 = E [X_t X_t']$$

If $m_0(i, j)$ is a element of M_0 matrix (size p x p) in the i^{th} row and j^{th} column,

$$m_0(i, j) = E [x(i, t) x(j, t)]$$

Expected value of a matrix is matrix of expected values of individual elements

$$m_0(i, j) = \frac{1}{n} \sum_{t=1}^n x(i, t) x(j, t)$$

n is no. of time periods

Multivariate Stochastic models

$$m_0(i, j) = \frac{1}{n} \sum_{t=1}^n \left(\frac{Q_{i,t} - \bar{Q}_i}{s_i} \right) \left(\frac{Q_{j,t} - \bar{Q}_j}{s_j} \right)$$

Q is the original random variable before standardization e.g., stream flow

i.e., $m_0(i, j)$ is correlation coefficient between the data at sites i and j at time t .

Therefore M_0 is the cross-correlation matrix of lag zero

Multivariate Stochastic models

- The expectation of $X_t X_{t-1}'$ is denoted by M_1

$$M_1 = E [X_t X_{t-1}']$$

If $m_1(i, j)$ is a element of M_1 matrix (size $p \times p$) in the i^{th} row and j^{th} column,

$$m_1(i, j) = E [x(i, t) x(j, t - 1)]$$

Expected value of a matrix is matrix of expected values of individual elements

$$m_1(i, j) = \frac{1}{n-1} \sum_{t=2}^n x(i, t) x(j, t - 1)$$

n is no. of time periods

Multivariate Stochastic models

$$m_1(i, j) = \frac{1}{n-1} \sum_{t=2}^n \left(\frac{Q_{i,t} - \bar{Q}_i}{s_i} \right) \left(\frac{Q_{j,t-1} - \bar{Q}_j}{s_j} \right)$$

Q is the original random variable before standardization e.g., stream flow

i.e., $m_1(i, j)$ represents lag one cross correlation between the data at sites i and j .

Therefore M_1 is the cross-correlation matrix of lag one.

Multivariate Stochastic models

Considering the model,

$$X_{t+1} = AX_t + B\mathcal{E}_{t+1}$$

Post multiplying with X_t' on both sides and taking the expectation, .

$$E[X_{t+1}X_t'] = AE[X_tX_t'] + \underbrace{BE[\mathcal{E}_{t+1}X_t']}_0$$

$$M_1 = AM_0 + 0$$

$$A = M_1M_0^{-1}$$

Multivariate Stochastic models

Post multiplying with X_{t+1}' on both sides and taking the expectation, ..

$$X_{t+1} = AX_t + B\epsilon_{t+1}$$

$$\underbrace{E[X_{t+1}X_{t+1}']}_{M_0} = AE[X_tX_{t+1}'] + BE[\epsilon_{t+1}X_{t+1}']$$

Multivariate Stochastic models

$$M_1 = E [X_t X_{t-1}']$$

$$M_1' = \left\{ E [X_t X_{t-1}'] \right\}'$$

$$= E \left[\left\{ X_t X_{t-1}' \right\}' \right]$$

$$= E [X_{t-1} X_t']$$

or

$$M_1' = E [X_t X_{t+1}']$$

Multivariate Stochastic models

$$\begin{aligned}\boldsymbol{\varepsilon}_{t+1}' X_{t+1}' &= \boldsymbol{\varepsilon}_{t+1}' \left\{ A X_t + B \boldsymbol{\varepsilon}_{t+1} \right\}' \\ &= \boldsymbol{\varepsilon}_{t+1}' X_t' A' + \boldsymbol{\varepsilon}_{t+1}' \boldsymbol{\varepsilon}_{t+1}' B'\end{aligned}$$

Taking expectation on both sides,

$$\begin{aligned}E \left[\boldsymbol{\varepsilon}_{t+1}' X_{t+1}' \right] &= E \left[\boldsymbol{\varepsilon}_{t+1}' X_t' A' + \boldsymbol{\varepsilon}_{t+1}' \boldsymbol{\varepsilon}_{t+1}' B' \right] \\ &= E \left[\boldsymbol{\varepsilon}_{t+1}' X_t' A' \right] + E \left[\boldsymbol{\varepsilon}_{t+1}' \boldsymbol{\varepsilon}_{t+1}' B' \right] \\ &= 0 + I B' \\ &= B'\end{aligned}$$

Multivariate Stochastic models

Substituting in the equation,

$$E \left[X_{t+1} X_{t+1}' \right] = AE \left[X_t X_{t+1}' \right] + BE \left[\varepsilon_{t+1} X_{t+1}' \right]$$

$$M_0 = AM_1' + BB'$$

$$M_0 = M_1 M_0^{-1} M_1' + BB'$$

$$BB' = M_0 - M_1 M_0^{-1} M_1'$$

$$A = M_1 M_0^{-1}$$

If $C = BB'$

$$C = M_0 - M_1 M_0^{-1} M_1'$$

Multivariate Stochastic models

- B does not have a unique solution.
- One method is to assume B is a lower triangular matrix.

$$BB' = \begin{bmatrix} b(1,1) & 0 & 0 & \cdot & \cdot & 0 \\ b(2,1) & b(2,2) & 0 & \cdot & \cdot & 0 \\ \cdot & \cdot & & & & \\ \cdot & \cdot & & & & \\ b(p,1) & b(p,2) & \cdot & \cdot & \cdot & b(p,p) \end{bmatrix} \begin{bmatrix} b(1,1) & b(2,1) & \cdot & \cdot & \cdot & b(p,1) \\ 0 & b(2,2) & \cdot & \cdot & \cdot & b(p,2) \\ \cdot & \cdot & \cdot & \cdot & \cdot & \cdot \\ \cdot & \cdot & \cdot & \cdot & \cdot & \cdot \\ 0 & 0 & \cdot & \cdot & \cdot & b(p,p) \end{bmatrix}$$

$$C = \begin{bmatrix} c(1,1) & c(1,2) & c(1,3) & \cdot & \cdot & c(1,p) \\ c(2,1) & c(2,2) & c(2,3) & \cdot & \cdot & c(2,p) \\ \cdot & \cdot & \cdot & \cdot & \cdot & \cdot \\ \cdot & \cdot & \cdot & \cdot & \cdot & \cdot \\ c(p,1) & c(p,2) & \cdot & \cdot & \cdot & c(p,p) \end{bmatrix}$$

Multivariate Stochastic models

- The diagonal elements of the B matrix are obtained as,

$$b(1,1) = c(1,1)^{1/2}$$

$$b(2,2) = \{c(2,2) - b^2(2,1)\}^{1/2}$$

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$$b(k,k) = \{c(k,k) - b^2(k,k-1) - b^2(k,k-2) - \dots - b^2(k,1)\}^{1/2}$$

Multivariate Stochastic models

- The elements in the k^{th} row are obtained as,

$$b(k,1) = \frac{c(k,1)}{b(1,1)}$$

$$b(k,2) = \frac{c(k,2) - b(2,1)b(k,1)}{b(2,2)}$$

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$$b(k,j) = \frac{c(k,j) - b(j,1)b(k,1) - b(j,2)b(k,2) \dots b(j,j-1)b(k,j-1)}{b(j,j)}$$

Multivariate Stochastic models

- Assumption is that the model is multivariate normal.

Example – 1

The annual flow in MCM at two sites X and Y is given below. Generate the first two values of data from these two sites.

Year	1	2	3	4	5	6	7	8	9	10
Annual flow at site X (MCM)	4946	7017	6653	6355	5908	5327	4548	3556	3852	5319
Annual flow at site Y (MCM)	5142	6240	5648	5977	6008	5045	4630	4604	4250	6182

Year	11	12	13	14	15	16	17	18	19
Annual flow at site X (MCM)	4631	5746	5111	5419	6060	7336	3736	3780	6034
Annual flow at site Y (MCM)	4703	6582	5461	5288	5440	7546	4634	4823	5577

Example – 1 (Contd.)

Site	X	Y
Mean	5333	5462
Std.dev.	1125.1	823.5

M_0 matrix is cross covariance matrix of lag zero

$$M_0 = \begin{matrix} & X & Y \\ \begin{matrix} X \\ Y \end{matrix} & \begin{bmatrix} r_{X,X}(0) & r_{X,Y}(0) \\ r_{Y,X}(0) & r_{Y,Y}(0) \end{bmatrix} \end{matrix}$$

Example – 1 (Contd.)

$$r_{X,Y}(0) = \frac{\sum_{i=1}^n (x_{X,i} - \bar{x}_X)(x_{Y,i} - \bar{x}_Y)}{(n)s_X s_Y}$$

$$M_0 = \begin{bmatrix} 1 & 0.796 \\ 0.796 & 1 \end{bmatrix}$$

M_1 matrix is lag one cross covariance matrix

$$M_1 = \begin{matrix} & X & Y \\ \begin{matrix} X \\ Y \end{matrix} & \begin{bmatrix} r_{X,X}(1) & r_{X,Y}(1) \\ r_{Y,X}(1) & r_{Y,Y}(1) \end{bmatrix} \end{matrix}$$

Example – 1 (Contd.)

$$r_{X,Y}(1) = \frac{\sum_{i=1}^n (x_{X,i} - \bar{x}_X)(x_{Y,i+1} - \bar{x}_Y)}{(n-1)s_X s_Y}$$

$$M_1 = \begin{bmatrix} 0.302 & 0.164 \\ 0.02 & -0.118 \end{bmatrix}$$

$$A = M_1 M_0^{-1} \quad M_0^{-1} = \begin{bmatrix} 2.73 & -2.17 \\ -2.17 & 2.73 \end{bmatrix}$$

Example – 1 (Contd.)

$$A = \begin{bmatrix} 0.47 & -0.21 \\ 0.31 & -0.37 \end{bmatrix}$$

$$C = M_0 - M_1 M_0^{-1} M_1'$$

$$C = \begin{bmatrix} 0.89 & 0.76 \\ 0.76 & 0.95 \end{bmatrix}$$

Example – 1 (Contd.)

$$c(1,1) = 0.89, c(1,2) = c(2,1) = 0.76,$$

$$c(2,2) = 0.95$$

$$b(1,1) = c(1,1)^{1/2} = (0.89)^{1/2} = 0.94$$

$$b(2,1) = \frac{c(2,1)}{b(1,1)} = \frac{0.76}{0.94} = 0.81$$

$$\begin{aligned} b(2,2) &= \left\{ c(2,2) - b^2(2,1) \right\}^{1/2} \\ &= \left\{ 0.95 - 0.81^2 \right\}^{1/2} = 0.54 \end{aligned}$$

Example – 1 (Contd.)

$$b(1,1) = 0.94, b(2,1) = 0.81, b(2,2) = 0.54$$

$$B = \begin{bmatrix} 0.94 & 0 \\ 0.81 & 0.54 \end{bmatrix}$$

$$X_{t+1} = AX_t + B\mathcal{E}_{t+1}$$

$$\begin{bmatrix} x_{X,t+1} \\ x_{Y,t+1} \end{bmatrix} = \begin{bmatrix} 0.47 & -0.21 \\ 0.31 & -0.37 \end{bmatrix} \begin{bmatrix} x_{X,t} \\ x_{Y,t} \end{bmatrix} + \begin{bmatrix} 0.94 & 0 \\ 0.81 & 0.54 \end{bmatrix} \begin{bmatrix} \mathcal{E}_{X,t+1} \\ \mathcal{E}_{Y,t+1} \end{bmatrix}$$

Example – 1 (Contd.)

$$\begin{bmatrix} x_{X,1} \\ x_{Y,1} \end{bmatrix} = \begin{bmatrix} 0.47 & -0.21 \\ 0.31 & -0.37 \end{bmatrix} \begin{bmatrix} x_{X,0} \\ x_{Y,0} \end{bmatrix} + \begin{bmatrix} 0.94 & 0 \\ 0.81 & 0.54 \end{bmatrix} \begin{bmatrix} \varepsilon_{X,1} \\ \varepsilon_{Y,1} \end{bmatrix}$$

The initial value of $x_{X,0}$ and $x_{Y,0}$ are considered as zero
 $\varepsilon_{X,1} = -0.134$ and $\varepsilon_{Y,1} = -0.268$

$$\begin{aligned} \begin{bmatrix} x_{X,1} \\ x_{Y,1} \end{bmatrix} &= \begin{bmatrix} 0.94 & 0 \\ 0.81 & 0.54 \end{bmatrix} \begin{bmatrix} -0.134 \\ -0.268 \end{bmatrix} \\ &= \begin{bmatrix} -0.126 \\ -0.254 \end{bmatrix} \end{aligned}$$

Example – 1 (Contd.)

$$\begin{bmatrix} x_{X,2} \\ x_{Y,1} \end{bmatrix} = \begin{bmatrix} 0.47 & -0.21 \\ 0.31 & -0.37 \end{bmatrix} \begin{bmatrix} x_{X,1} \\ x_{Y,1} \end{bmatrix} + \begin{bmatrix} 0.94 & 0 \\ 0.81 & 0.54 \end{bmatrix} \begin{bmatrix} \varepsilon_{X,2} \\ \varepsilon_{Y,2} \end{bmatrix}$$

$$\varepsilon_{X,2} = 1.639 \text{ and } \varepsilon_{Y,2} = 0.134$$

$$\begin{aligned} \begin{bmatrix} x_{X,2} \\ x_{Y,1} \end{bmatrix} &= \begin{bmatrix} 0.47 & -0.21 \\ 0.31 & -0.37 \end{bmatrix} \begin{bmatrix} -0.126 \\ -0.254 \end{bmatrix} + \begin{bmatrix} 0.94 & 0 \\ 0.81 & 0.54 \end{bmatrix} \begin{bmatrix} 1.639 \\ 0.134 \end{bmatrix} \\ &= \begin{bmatrix} 1.543 \\ 1.449 \end{bmatrix} \end{aligned}$$

Example – 1 (Contd.)

Generated annual flow at Site X :

Site	X	Y
Mean	5333	5462
Std.dev.	1125.1	823.5

$$Q_{X,1} = \bar{x}_X + x(X,1)s_X$$

$$Q_{X,1} = 5333 - 0.126*1125.1 = 5191.2 \text{ MCM}$$

$$Q_{X,2} = \bar{x}_X + x(X,2)s_X$$

$$Q_{X,2} = 5333 + 1.543*1125.1 = 7069 \text{ MCM}$$

Similarly at Site Y :

$$Q_{Y,1} = 5462 - 0.254*823.5 = 5252.8 \text{ MCM}$$

$$Q_{Y,2} = 5462 + 1.449*823.5 = 6655.3 \text{ MCM}$$

CASE STUDIES